# Changes in the spatial distribution of fishes in lakes along a residential development gradient 

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#### Abstract

As the human demand for freshwater natural resources such as fish and drinking water increases, we may rely more heavily on models to predict the response of aquatic ecosystems to natural and anthropogenic disturbance. Theses models in turn implicitly depend on the underlying spatial distribution of organisms. In terrestrial ecosystems, increased natural resource utilization has transformed habitat and changed the spatial distribution of organisms, with subsequent negative effects on biota. Recent studies in lakes demonstrate that human development of lakeshores alters the physical habitat and nutrient cycles. The impact of such disturbance by humans on the spatial distribution of aquatic organisms however, remains unknown. Here we quantify the effect of lakeshore development on the spatial distribution of fishes in 23 lakes in the U.S. Pacific Northwest. We found a significant decrease in the spatial aggregation of fishes with increased shoreline development by humans, reflecting a loss of refugia and resource heterogeneity that favor aggregation among fishes. We also found that lakes with a high perimeter to surface area ratio and a relatively shallow littoral zone had much higher levels of fish aggregation, suggesting the importance of terrestrial inputs to lakes. Finally, we found a marginally significant decrease in fish spatial aggregation with increased total phosphorus concentration, but no effect of chlorophyll concentration, water transparency, the predator;prey ratio, or number of species on fish spatial distributions. These results suggest that anthropogenic modification of shorelines is significantly altering the spatial distribution of important aquatic organisms, and that these changes may have important implications for predictive modeling of ecosystem dynamics.


Key words: spatial distribution, lakeshore development, morphometry, model, hydroacoustics

## INTRODUCTION

Humans increasingly impact the world's ecosystems through a variety of disturbances such as changing land use, species introductions and extinctions, alteration of nutrient cycles, and the use of $>50 \%$ of the available fresh water (Vitousek et al. 1997). As the global human population grows we will rely more heavily on aquatic resources for food, drinking water, and recreation (Naiman et al. 1998, Naiman and Turner 2000, Jackson et al. 2001b). In light of these increased demands, ecologists are challenged with making predictions about the response of aquatic ecosystems to changes in the abiotic and biotic environment under a variety of natural and anthropogenic disturbances (Pace 2001, Carpenter 2002). Central to these predictions are models that incorporate predator feeding and prey mortality rates; these models depend implicitly on the spatial distributions of interacting organisms. Furthermore, for simplicity these models often assume that there is no systematic change in the spatial distribution of organisms with human disturbance. Nevertheless, this critical assumption remains largely untested for aquatic systems (Boisclair 2001), despite some evidence showing that human-mediated effects can alter habitat use by fishes with cascading effects on lower trophic levels (Jeppesen et al. 2000, Olin et al. 2002). Furthermore, in terrestrial systems, anthropogenic habitat loss and fragmentation has altered the spatial distribution of organisms, and changed competitive and predator-prey interactions (Debinski and Holt 2000, Haila 2002).

Lakeshores throughout North America and Europe are increasingly being developed and altered by humans (Ostendorp et al. 1995, Christensen et al. 1996, Jennings et al. 1999, Radomski and Goeman 2001, Beeton 2002, Gulati and van Donk 2002). Accompanying this development are anthropogenic stresses to lake ecosystems including disruption of nutrient cycles resulting in eutrophication (Jeppesen et al. 1997, Carpenter et al. 1998, Garrison and

Wakeman 2000, Moore et al. 2003) and changes in fish community composition (Jeppesen et al. 2000, Olin et al. 2002), introduction of exotic species (Jackson et al. 2001a), and the exploitation of predatory fishes by recreational and commercial fisheries (Pauly et al. 2001, Post et al. 2002). Furthermore, transformation of nearshore habitats adversely affects fish communities because they form part of the physical template on which lake ecosystems lie (Jackson et al. 2001a, Olden and Jackson 2001, Olden et al. 2001). Some direct impacts of lakeshore residential development on the physical structure of lake habitats include bottom substrate modifications (Beauchamp et al. 1994) and the removal of coarse woody debris (Christensen et al. 1996) and aquatic vegetation (Ostendorp et al. 1995, Radomski and Goeman 2001). These habitat modifications have been associated with a reduction of fish growth rates (Schindler et al. 2000), altered trophic interactions (Jeppesen et al. 1997), and have the potential to alter fish spatial distributions (Schindler and Scheuerell 2002).

Detecting changes in fish assemblages caused by human disturbance usually requires observations before and after the disturbance (Lester et al. 1996), but such studies are often protracted, expensive, and confounded by multiple time-dependent processes. We took an alternative, comparative approach across lakes, which is useful in ecosystem studies where experimental replication is impossible (Schindler 1998) and more rapid feedback is desired. To evaluate the cumulative effect of lakeshore development on the spatial distribution of fishes, we sampled 23 lakes in western Washington, USA, and southwestern British Columbia, Canada. These lakes spanned a residential development gradient in the greater Seattle, WA metropolitan area, which has a current human population near 2 million that grew almost $20 \%$ in the last 10 years. The intensity of lake shoreline disturbance by humans is highest near the Seattle urban center and then generally decreases with distance from the urban area. Here we explored whether
human disturbance to lakeshores has any consequences for the spatial distribution of fishes. We also examined the potential role of the lakes' physical morphology on fish spatial distributions because of the natural interaction between morphology and physical habitat available to aquatic organisms. Lakes with very steep sides have relatively little littoral habitat compared to pelagic habitat (Wetzel 2001). Lakes with high perimeter to area ratios can receive relatively high inputs of terrestrial matter compared to more circular lakes (Wetzel 2001). Finally, we compared the relative role of cultural eutrophication (water clarity, total phosphorus and chlorophyll concentrations), species diversity, and the ratio of predators to prey in explaining the observed variance in fish spatial distributions.

## METHODS

For our analyses, we considered two multiple regression models to explain the observed variation in fish spatial distributions. Using backward stepwise regression, we began with the following 8 predictor variables: housing density, shoreline morphometry $\left(D_{\mathrm{L}}\right)$, basin morphometry $\left(D_{\mathrm{V}}\right)$, total phosphorus concentration, chlorophyll concentration, extinction coefficient ( $k$ ), the predator:prey ratio, and the number of species. Model 1 uses the number of houses per km of shoreline whereas Model 2 uses the number of houses per ha of lake surface area.

Lake data. We sampled 10 of the lakes during 11-20 August 1998 and the other 13 during 23-31 August 1999. All of the study lakes were located in western Washington, USA, and southwestern British Columbia, Canada, in a region east of Puget Sound and west of the Cascade Mountains in an attempt to control for differences in geology, elevation, lake size, and morphometry. Information on lake surface area, shoreline length, mean depth, and maximum
depth was obtained from the Washington State Lake Survey Data (Bortleson et al. 1976) and the British Columbia Ministry of Sustainable Resource Management (http://www.gov.bc.ca/srm). The elevation of all lakes ranged from $35-520 \mathrm{~m}$ above sea level. The average lake surface area was 22 ha ( $\pm 3.4 \mathrm{SE})$, the mean depth across all lakes was $8.6 \mathrm{~m}( \pm 1.1 \mathrm{SE})$, and the average maximum depth was $18 \mathrm{~m}( \pm 2.5 \mathrm{SE})$. All of the WA lakes had public access for boating and fishing, but the BC lakes were located at the University of British Columbia Malcolm Knapp Research Forest and therefore had neither public access nor shoreline development.

We used the number of nearshore houses per km of shoreline and the number of nearshore houses per ha of lake area as general indices for the intensity of human disturbance. We also chose three specific indicators of cultural eutrophication as potential predictors: 1) hypolimnetic total phosphorus (TP) concentration ( $\mu \mathrm{g} \mathrm{L}^{-1}$ ), 2) mean epi- and metalimnetic chlorophyll a concentration $\left(\mu \mathrm{g} \mathrm{L}^{-1}\right)$, and 3 ) the light extinction coefficient $\left(k, \mathrm{~m}^{-1}\right)$. Lake water was collected from depth using an 8 L Van Dorn bottle. Following persulfate digestion and subsequent reaction with molybdate and stannous chloride, TP concentration was determined colorometrically by reading the samples on a spectrophotometer at 720 nm and comparing them to a standard curve (Anonymous 1999). Chlorophyll $a$ concentration was determined by filtering lake water through $0.7 \mu \mathrm{~m}$ GF/F glass-fiber filters, freezing them, extracting with methanol, and then reading on fluorometer and correcting for pheophytins (Marker et al. 1980). Light was measured at $1-\mathrm{m}$ intervals using a Li-Cor Model 192SA light sensor. For each lake, we fit $k$ from the standard equation for light extinction, where $L_{z}=L_{0} \exp [-k z], L_{z}$ is the amount of light at a given depth $z$, and $L_{0}$ is the amount of light at the lake surface.

We chose two simple metrics to describe the physical morphology of our study lakes. First, we used the volume development index $\left(D_{\mathrm{V}}\right)$ as a measure of the lake basin morphometry.

The index is expressed as the ratio of a lake's mean depth to its maximum depth (Wetzel 2001), such that for a conical basin $D_{\mathrm{V}}=0.33$ and for a perfectly cylindrical basin $D_{\mathrm{V}}=1.0$. Second, we used the shoreline development index $\left(D_{\mathrm{L}}\right)$ to classify the shape of the lake shoreline relative to its area, where $D_{\mathrm{L}}=0.5 P(\pi A)^{-0.5}, P$ is the length of shoreline, and $A$ is the lake surface area (Wetzel 2001). For these purposes we adopt the convention of referring to $D_{\mathrm{L}}$ as the shoreline morphometry index to avoid confusion with our classification of the intensity of human shoreline development.

Fish data. We obtained direct estimates of the fish community composition in each lake with 10-12 unbaited minnow traps and 2 variable-mesh gill nets, set for approximately 2 hours. The minnow traps had 0.25 inch mesh and were submerged in $0.5-2 \mathrm{~m}$ of water in the littoral zone. Each gillnet was $33 \times 2 \mathrm{~m}$ and consisted of five equal-area panels with bar mesh sizes of $1.3,2.5,3.8,5.1$, and 6.4 cm , and was set perpendicular to shore along the bottom. We gathered additional information on the fish community composition from Washington Department of Fish and Wildlife stocking and sampling records (http://www.wa.gov/wdfw) due to our low CPUE presumably resulting from low fish densities and relatively short soak times. The fish community was similar among all lakes and consisted of naturally reproducing and stocked populations of 15 of the following species: largemouth bass (Micropterus salmoides), pumpkinseed sunfish (Lepomis gibbosus), black crappie (Pomoxis nigromaculatus), yellow perch (Perca flavescens), brown bullhead (Ictalurus nebulosus), rainbow trout (Oncorhynchus mykiss), or cutthroat trout (O. clarkii).

We measured the three-dimensional locations of fish using a BioSonics DT6000 splitbeam echosounder operating at 200 kHz with an elliptical transducer consisting of -3 dB beam angles of 6.7 and 13.5 degrees; the narrow beam was parallel to the direction of travel. We used
a 0.4 ms pulse at a rate of $4 \mathrm{~s}^{-1}$ with the lower threshold set at -70 dB . Using the BioSonics Visual Acquisition 4.0 software, all incoming signals were digitized at the transducer and then stored on a computer hard drive for later analysis. The transducer was suspended from a pole mount immediately under the water surface and towed alongside the boat at a speed of $10 \mathrm{~km} \mathrm{hr}^{-}$ ${ }^{1}$ as the survey followed 8-14 predetermined transects across each lake over a total distance of $1300-4200 \mathrm{~m}$ (mean $=2500 \pm 130 \mathrm{SE})$. Acoustic transects were selected non-randomly in an effort to provide adequate survey coverage without biasing any particular depth contours, embayments, docks, houses, etc. The total sample volume was $5800-160000 \mathrm{~m}^{3}$ (mean $=30000$ $\pm 8300$ SE). All hydroacoustic surveys occurred between 11:30 and 17:30. Fish density and the depth and size of individual echoes were obtained using the BioSonics DT Analyzer 4.0 software, but we only analyzed depths <1 m due to near-field effects of the acoustic beam. For these analyses we combined all potential prey fish together and classified all targets less than -46 dB as prey fish, which correspond to fish lengths < 100 mm (Burczynski and Johnson 1986). All subsequent analyses only refer to these small fishes. Densities of predator and prey fishes were obtained by multiplying the total fish density by the proportion of all target strength values within each class (prey <-46 dB; predators > -40 dB ). We obtained simultaneous measurements on the $x-y$ coordinates of the fish by coupling the hydroacoustics to a sub-meter accurate Trimble ${ }^{\circledR}$ AgGPS 132 differential global positioning system (DGPS).

We assessed the degree of spatial aggregation among fishes using a method developed for three-dimensional point process patterns (Scheuerell 2002) based on the methodology for a two-dimensional case (Coomes et al. 1999). This method involves measuring the locations of individual organisms in three dimensions with hydroacoustics coupled to a differential global positioning system (DGPS), and then calculating the nearest neighbor distances among
organisms. For each lake, we compared the observed cumulative probability ( $G_{\text {obs }}$ ) of a fish having a nearest neighbor within a given radius to that expected in a random distribution ( $G_{\mathrm{PV}}$ ) assuming a three-dimensional Poisson process. We compared the two probability distributions through the test statistic $d w_{\mathrm{obs}}$, which equals the maximum difference between $G_{\mathrm{obs}}$ and $G_{\mathrm{PV}}$ over all radii examined ( $0.2-5.0 \mathrm{~m}$ ). This statistic therefore tests for aggregation over all spatial scales within the range of radii considered (Coomes et al. 1999). We estimated the null distribution of $d w$ by comparing $G_{\mathrm{PV}}$ to $G_{\mathrm{obs}}$ from 1000 Monte Carlo simulations of a threedimensional Poisson process. The critical values of $d w$ were then calculated from the upper tail of the null distribution such that the $95^{\text {th }}$ percentile indicated $d w_{\text {crit }}$ for $P=0.05$ (sensu Coomes et al. 1999). When the ratio of $d w_{\text {obs }}: d w_{\text {crit }}>1$, the spatial distribution is significantly aggregated. We used this ratio as a measure for the intensity of aggregation among fishes such that large values indicate strong aggregation and values approaching zero are randomly distributed in space. To account for sampling biases associated with the beam geometry, we distributed points randomly within a large volume according to a three-dimensional Poisson process and then sampled them using an artificial hydroacoustics beam using 500 Monte Carlo simulations. We then calculated the edge correction as the mean difference between the observed $G_{\mathrm{PV}}$ and the expected $G_{\mathrm{PV}}$ in an infinite universe (Coomes et al. 1999).

## RESULTS AND DISCUSSION

Both human development of lakeshores and lake morphometry were related to the spatial distribution of fishes (Table 1). In the stepwise multiple regression, housing density had the strongest effect on the spatial aggregation of fishes, such that as housing density increased, the spatial aggregation of fishes decreased (Figure 1A, 1B). This pattern held for either case when
we used the number of houses per km of shoreline or the number of houses per ha of lake surface area, although the overall effect of houses per area was slightly more than that of houses per shoreline (Table 1). Furthermore, in both cases the basin morphometry index $\left(D_{\mathrm{V}}\right)$ had the next strongest impact (Table 1). $D_{\mathrm{V}}$ had a negative effect on the degree of spatial aggregation among fishes (Figure 1D), indicating that as the lake basin became more cylindrical in shape, fish aggregation decreased. Once the ratio of mean depth to maximum depth reached 0.5 , the fish showed essentially no aggregation behavior. We also found a significant positive effect of lakeshore morphometry $\left(D_{\mathrm{L}}\right)$ on the degree of spatial aggregation among fishes (Figure 1C), demonstrating that fish were more aggregated in lakes with increasingly complex shorelines. However, $D_{\mathrm{L}}$ was only a significant predictor in the case where we used the number of houses per shoreline, and in that case only marginally so (Table 1). Lakes with a high perimeter to area ratio and a relatively shallow littoral zone may offer more habitat heterogeneity and refuge from predation, which could subsequently affect the spatial distribution of fishes (Olden and Jackson 2001).

Of the cultural eutrophication variables we examined, only total phosphorus had a significant effect on the degree of spatial aggregation among fishes, but it was relatively weak and only occurred when we used houses per area as the metric for human disturbance (Model 2, Table 1). The trend was for a decrease in fish spatial aggregation with increased TP concentration (Figure 1E). Although it was not significant, increased chlorophyll concentration also showed a pattern toward decreased aggregation (Figure 1F). This could suggest that decreased water clarity associated with increased nutrients and algae also decreased fish aggregation, as fish spatial distributions are often affected by light (Sogard and Olla 1993, Miyazaki et al. 2000, Scheuerell and Schindler 2003). However, we found no relationship
between our direct measure of water clarity $(k)$ and fish spatial aggregation (Figure 1G). On the other hand, it might suggest that changes in the lower food web structure associated with increased anthropogenic activities alter the spatial distribution of fishes through some sort of bottom-up forcing associated with shifts in fish diets from predominantly benthic in undeveloped areas to more pelagic in heavily impacted areas (Schindler and Scheuerell 2002). Predators often exhibit strong effects on the spatial distribution of fishes (Werner and Gilliam 1984, Lima and Dill 1990, Lima 1998). However, we failed to find any statistical relationship between the ratio of predators to prey and the degree of spatial aggregation in our study lakes (Table 1, Figure 1H). We also found no effect of the actual predator density on spatial aggregation (linear regression, $\left.R^{2}=0.094, P>0.10\right)$.

The density of organisms is implicitly coupled to their spatial distribution; as density increases, the expected distance between neighbor's decreases. The average density of prey fishes in the study lakes was $0.097 \mathrm{~m}^{-3}( \pm 0.024 \mathrm{SE})$ and the mean density of predatory fishes was $0.0011 \mathrm{~m}^{-3}( \pm 0.00038 \mathrm{SE})$. Neither the density of prey fishes nor predatory fishes varied across the residential shoreline development gradient (linear regression, prey: $R^{2}=0.005, P>$ 0.10; predators: $\left.R^{2}=0.0004, P>0.10\right)$. It is important to point out that the spatial statistic we used ( $d w$ ) explicitly accounts for the density of organisms, and is therefore robust to potential biases associated with differences in fish density among lakes. We also found no significant change in the mean acoustic target strength (i.e. fish size) with residential development (linear regression, $\left.R^{2}=0.15, P>0.10\right)$. The mean $( \pm \mathrm{SE})$ number of fish caught per lake was $6.3 \pm 0.83$ and $4.0 \pm 0.49$ in the gillnets and minnow traps respectively, which was generally too low to assign reasonable species identification to the acoustic density estimates. Therefore, we used the
trap and net data to assess species presence and community composition, and relied on the hydroacoustics for fish density.

Human disturbance has been associated with changes in fish community composition in streams (Vila-Gispert et al. 2002), wetlands (Brazner 1997), lakes (Jeppesen et al. 2000), and nearshore marine environments (Guidetti et al. 2002). Differences in community composition could affect the spatial distribution of fishes through predator-prey interactions (Brabrand and Faafeng 1993) or other behavioral mechanisms (Olden and Jackson 2001). Although there were subtle differences in the fish community composition among our study lakes (Figure 2), we found no effect of the number of species on the spatial aggregation of fishes in a stepwise multiple regression context (Table 1) or when we considered species number by itself (Pearson correlation, $r=0.062, P>0.10$ ). Furthermore, we found no significant effects of the presence or absence of any particular fish species on the spatial distribution of fishes (Kruskal-Wallis test, $H$ $=4.2, P>0.10$ ). Lastly, we found no difference in the mean number of fish species in lakes with significant aggregation versus lakes without (Mann-Whitney test, $U=91, P>0.10$ ). This suggests that the cumulative effects of humans and lake morphometry are more important than the actual fish species assemblage in lakes such as these with relatively simple fish communities.

Our sampling design and methodology was somewhat biased toward the pelagic habitat and deeper fish due to the cross-lake transects and the conical shape of the hydroacoustic beam. However, there should not have been any systematic bias against any particular lake because transects did not follow any specific depth contour, the length of which would differ among lakes. Furthermore, we also note that our analysis of the sampled fish explicitly accounts for the beam geometry and therefore doesn't introduce any bias associated with fish depth. Given the
possible bias toward pelagic habitats, our results using the number of houses per ha of lake surface area may be more appropriate (Table 1, Model 2).

Our comparisons relied on relatively simple, linear models to explain differences in the spatial distribution of fishes. Certainly the ecological processes occurring throughout the lakes vary spatially and are affected by a variety of internal and external processes (Turner and Carpenter 1999). This observed variation may result from nonlinear interactions among ecosystem components, such as predators and prey (De Roos and Persson 2002). Largemouth bass were the dominant piscivore in our lakes, occurring in 15 of 23 lakes across a wide range of human development intensity (Figure 2). Largemouth bass tend to be highly aggregated in undeveloped lakes (Essington and Kitchell 1999), but that appears to change as lakes are impacted by humans. Furthermore, human-mediated nutrient loading, habitat disturbance, and exploitation can abruptly shift lakes to alternative stable states (Scheffer et al. 2001). Finally, habitats within lakes are not isolated, but are instead connected via biotic and abiotic processes, which are likely a function of the physical template (Schindler and Scheuerell 2002). Perhaps our observed change in the spatial distribution of prey fishes resulted from nonlinear interactions among their predators, human impacts, and lake morphometry.

All natural populations demonstrate spatial and temporal variation in their numbers.
Some of this variability is associated with the physical structure of their habitat. For many years terrestrial ecologists have paid particular attention to the role of human-mediated habitat fragmentation in determining the distribution and interaction among species. These studies have shown that changes in habitat structure often lead to subsequent variation in the spatial distribution of organisms and that these changes must be considered when conserving or restoring ecosystems (Debinski and Holt 2000, Fahrig 2002, Haila 2002). Only recently have we
begun to appreciate analogous anthropogenic disturbances in aquatic ecosystems. Evidence is accumulating which shows that human development of lakeshores affects the physical habitat of lakes and the fish community associated with it (Beauchamp et al. 1994, Christensen et al. 1996, Jennings et al. 1999, Schindler et al. 2000, Jackson et al. 2001a, Olden and Jackson 2001, Radomski and Goeman 2001). This study represents another step toward understanding the negative effects of human disturbance on lake ecosystems by showing how the spatial distribution of fishes changes with increased lakeshore development. We suspect that a combination of eutrophication and associated water clarity (Jeppesen et al. 2000, Olin et al. 2002), changes in trophic structure associated with fish exploitation (Post et al. 2002) and predation (Brabrand and Faafeng 1993), and degradation of nearshore habitats (Brazner 1997, Schindler and Scheuerell 2002) all contribute to this response. Furthermore, direct disturbance by swimmers and boaters could be perceived as a form of predation risk and alter fish spatial distributions (sensu Frid and Dill 2002).

Researchers and managers alike often rely on empirical models of fish production, nutrient cycling, or contaminant flows to predict the effect of human actions on lake ecosystems (Carpenter and Gunderson 2001). For simplicity, these models often ignore or average spatially explicit features of the biotic and abiotic environments. However, lakes are heterogeneous, threedimensional landscapes whose structure is influenced by physical, biological, and human activities (Riera et al. 2001). Therefore, consideration of the underlying assumptions is critical when a model is used to evaluate changes within or among ecosystems (Boisclair 2001). As an example we show how lakeshore development changes the underlying spatial distribution of fishes, which is usually assumed constant across systems. Most of the anthropogenic disturbances to aquatic ecosystems associated with lakeshore development occur gradually over
time (Jennings et al. 1999), which should allow for proactive management in areas where development of lakeshores is just beginning (Moore et al. 2003). In these cases, incorporating knowledge about the interaction between humans and the environment will prove critical to understanding and predicting ecosystem dynamics (Carpenter 2002). Given the wide-ranging negative effects of anthropogenic habitat transformations on organisms in terrestrial ecosystems, we suspect that similar effects are prevalent, but unexplored, in aquatic ecosystems.

## ACKNOWLEDGMENTS

We thank Michael Richland and Wendy Palen for assistance in the field and the University of Washington Royalty Research Fund for financial support. Shahid Naeem, Bob Paine, Wendy Palen, Erik Jeppesen, and two anonymous reviewers provided useful comments on earlier drafts of the manuscript.

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Table 1. Results of the backward-stepwise multiple regression analyses to predict the spatial aggregation of fishes. Both statistical models began with the following predictor variables: housing density, shoreline morphometry $\left(D_{\mathrm{L}}\right)$, basin morphometry $\left(D_{\mathrm{V}}\right)$, total phosphorus concentration, chlorophyll concentration, extinction coefficient $(k)$, the predator:prey ratio, and the number of species. Model 1 uses the number of houses per km of shoreline whereas Model 2 uses the number of houses per ha of lake surface area.

| Model | Variable | Coef. | SE | $t$ | $P$ | $r^{2}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | intercept | 3.1 | 1.8 | 1.8 | 0.096 |  |
|  | houses | -0.038 | 0.011 | -3.5 | 0.002 | 0.32 |
|  | basin | -6.9 | 2.3 | -3.0 | 0.007 | 0.24 |
|  | shoreline | 1.4 | 0.68 | 2.1 | 0.050 | 0.10 |
|  | overall |  |  |  | $<0.001$ | 0.66 |
| 2 | intercept | 6.7 | 1.1 | 6.4 | $<0.001$ |  |
|  | houses | -0.40 | 0.091 | -4.4 | $<0.001$ | 0.39 |
|  | basin | -9.1 | 2.1 | -4.4 | $<0.001$ | 0.24 |
|  | TP | -0.014 | 0.0080 | -1.8 | 0.088 | 0.09 |
|  | overall |  |  |  | $<0.001$ | 0.72 |

## FIGURE LEGENDS

Figure 1. Relationship between the spatial aggregation of fishes and A) houses per km of shoreline, B) houses per ha of lake surface area, C) shoreline morphometry index $D_{\mathrm{L}}$, D) lake basin morphometry index $D_{\mathrm{V}}, \mathrm{E}$ ) hypolimnetic total phosphorus concentration, F) mean epi- and metalimnetic chlorophyll $a$ concentration, G) light extinction coefficient $k$, and H ) the ratio of predators to prey (by density). The degree of spatial aggregation equals the ratio of $d w_{\text {obs }}$ to $d w_{\text {crit. }}$ Values $>1$ indicate significant aggregation at $\alpha=0.0022$ ( $=0.05 / 23$ comparisons). See methods for descriptions of $D_{\mathrm{L}}, D_{\mathrm{V}}$, and $k$.

Figure 2. The total number of fish species and community composition for the 23 study lakes ranked in order of increasing housing density (listed in parentheses, $\mathrm{km}^{-1}$ ). See methods for individual species names.




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